Estimating Reductions in Ethnic Inequalities in Child Adiposity from Hypothetical Diet, Screen Time, and Sports Participation Interventions

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Background: Childhood obesity is a global epidemic, and its prevalence differs by ethnicity. The objective of this study was to estimate the change in ethnic inequalities in child adiposity at age 10 resulting from interventions on diet at age 8 and screen time and sports participation at age 9.

Methods: We conducted a population-based cohort study, the Generation R Study, from 9,749 births in Rotterdam (2002–2006), of which 9,506 children remained in the analysis. We measured ethnicity, diet, screen time, and sports participation through questionnaires; we measured weight, body mass index (BMI), fat mass index, and fatfree mass index directly. We used sequential G-estimation to estimate the reduction in inequality that would result from the interventions.

Results: We observed that sociodemographic characteristics, diet, screen time, sports participation, and all adiposity measurements were more favorable in children from Western versus non-Western ethnic backgrounds: weight = -1.2 kg (95% confidence interval [CI] = -1.7, -0.8), BMI = -1.0 kg/m^2 (CI = -1.2, -0.9), and fat mass index = -0.8 kg/m^2 (CI = -0.9, -0.7). We estimated that extreme intervention (maximum diet score of 10, no screen time, and >4 hours/ week of sports) reduced ethnic inequalities by 21% (CI = 8%, 35%) for weight, 9% (CI = 4%, 14%) for BMI, and 9% (CI = 6%, 13%) for fat mass index. A diet score ≥ 5 points, screen time ≤ 2 hours/day, and sports participation >2 hours/week reduced ethnic inequalities by 17% (CI = 6%, 28%) for weight, 7% (CI = 3%, 11%) for BMI, and 7% (CI = 4%, 10%) for fat mass index.

Conclusions: Our results are consistent with the hypothesis that interventions integrating diet, screen time, and sports participation have a moderate impact on reducing ethnic inequalities in child adiposity.

Keywords: Adiposity; Body mass index; Child; Diet; Ethnic inequalities; Screen Time; Sports Participation

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besity is a global epidemic and public health problem in both adults and children. During the last 4 decades, its worldwide prevalence has nearly tripled in adults and increased 10-fold in children and adolescents. In 2016, 13% of adults and 7% of children and adolescents had obesity. 1,2

According to the World Health Organization, obesity is a major risk factor for cardiovascular diseases, diabetes, osteoarthritis, and some cancers.1 Previous research suggests that early prevention of obesity in children may be the best approach to prevent future increases in morbidity and healthcare costs that will likely occur as children with obesity age.3

Obesity is not equally distributed throughout the population. Ethnic minorities in European countries have been found to have a higher burden of overweight and obesity both in adults and children. 4-6 However, little is known about the degree to which specific factors are driving these inequalities and the pathways through which differences in body mass and body composition occur. Improving our knowledge of these mechanisms will inform where public health interventions can have the largest impact to reduce ethnic inequalities in obesity in the population from childhood onward.

The Netherlands has a long history of immigration.⁷ Currently, of its 17.3 million inhabitants, 24% are immigrants or the children of immigrant parents.8 The country has committed to reducing socioeconomic inequality in life expectancy by 25% by the year 2020 yet it appears that despite the implementation of various intervention programs, the country has not yet achieved this goal. In the period 2011–2014, life expectancy was 85 among those with the highest household income and 77 years among those with the lowest household income.¹⁰ Reducing ethnic inequalities in cardiovascular risk factors that mediate the relationship between child adiposity and mortality due to cardiovascular disease could help in achieving this commitment.11-14

The objective of this study was to estimate the change in ethnic inequalities in child adiposity at age 10 resulting from hypothetical interventions on diet at age 8 and screen time and sports participation at age 9. We used sequential G-estimation to separate the estimation of the inequality and the model for the interventions separately to avoid adjusting the inequality itself for covariates, which occurs when both models are fit together.

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METHODS

Study Design

We used data from the Generation R Study, a populationbased prospective cohort study whose purpose is to identify early environmental and genetic determinants of growth, development, and health from fetal life until young adulthood. 15

The Generation R Study is set in Rotterdam, the second largest city in the Netherlands. As a port and industrial city, its economic growth has led to large-scale migration.¹⁶ In 2019, Rotterdam had 644,618 inhabitants and 29% of them born abroad. The largest immigrant communities were from Suriname, Turkey, Morocco, and the Dutch Caribbean.¹⁷

Study Population

All pregnant women living in Rotterdam and with an expected delivery date between April 2002 and January 2006 were invited to participate. At baseline, 9,778 mothers were enrolled in the study of which 9,749 had live born children.¹⁵ A flow chart for this study is presented in the Figure.

Collection of Information **Ethnic Background**

A participant was considered to be of non-Dutch ethnic origin if one of their parents was born abroad as reported by their parents on a questionnaire applied at baseline. If both parents were born abroad, the country of birth of the participant's mother determined the child's ethnic background. 15 A dichotomous variable, 1: "Western ethnic background" and 0: "Non-Western ethnic background," was constructed according to the national classification by Statistics Netherlands.¹⁸ It was designed to reflect the social and economic position

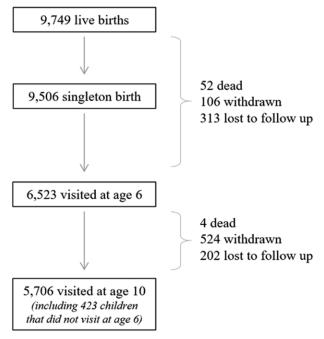


FIGURE. Flow chart of the participant included in this study.

of people in Dutch society. Western ethnicity included Dutch, European, North American, Oceanian, Japanese, and Indonesian (former Dutch East Indies). Non-Western ethnicity included Turkish, African, Latin-American, and Asian (except Indonesia and Japan).

Diet, Screen Time, and Sports Participation

A semiquantitative food frequency questionnaire was completed by the parents when the child was 8 years old. Frequency and amount of food consumed by the child during the last 4 weeks was registered and then converted into g/day. A diet quality score was computed, by calculating the ratio of reported and recommended food intake by age group for 10 items: fruit (≥150 g/d), vegetables (≥150 g/d), whole grains ($\geq 90 \,\mathrm{g/d}$), fish ($\geq 60 \,\mathrm{g/week}$), legumes ($\geq 84 \,\mathrm{g/week}$), nuts (≥15 g/d), dairy (≥300 g/d), oils and soft or liquid fats (≥30 g/d), sugar-containing beverages (≤150 g/d), and highfat and processed meat (≤250 g/week). 19 Total score for diet quality ranged from 0 to 10 on a continuous scale, with a higher score indicating a better overall diet quality. The methodology to obtain this diet score has been published previously.²⁰

Screen time and sports participation were assessed by parent-reported questionnaire when the child was 9 years old. Screen time included television viewing, video/DVD use, and computer game use. Frequency and duration of these activities were asked for weekdays and weekend days separately. We combined these variables to estimate the average daily screen time.21 Sports considered included soccer, hockey, basketball, handball, korfball, tennis, judo, karate, gymnastics, jazz ballet, etc. We did not include school sports activities such as physical education lessons and swimming lessons in this variable. Sports participation was categorized into less than 1 hour, from 1 to 2 hours, from more than 2 to 4 hours and more than 4 hours a week.

Adiposity Indicators

Child anthropometrics and body composition were measured at 6 and 10 years old by trained staff in a dedicated research center at Erasmus MC-Sophia Children's Hospital. Height was determined in standing position to the nearest millimeter without shoes with a Harpenden stadiometer (Holtain Limited, Dyfed, United Kingdom). Weight was measured using a mechanical personal scale (SECA, Almere, The Netherlands). Body mass index was calculated as body weight (kg)/height (m)². Body composition was measured using a dual-energy X-ray absorptiometry (DXA) scanner (iDXA; GE-Lunar, 2008, Madison, WI). Fat mass index [fat mass (kg)/height (m)²] and fat-free mass index [fat-free mass (kg)/ height $(m)^2$ were calculated.

This study was conducted in accordance with the guidelines proposed in the World Medical Association's Declaration of Helsinki and has been approved by the Medical Ethical Committee at Erasmus Medical Center, University Medical

Center Rotterdam. Written informed consent was obtained from all participants' parents.

Analysis

Recent thinking on health inequalities and finding interventions to reduce them has made it clear that, although these questions resemble mediation analyses, there are important differences.²²⁻²⁴ For example, the degree to which a health inequality would change if an intervention was implemented is not a type of controlled direct effect because we are not interested in intervening on the group that defines the inequality, ethnicity, in this case.²⁴ Our parameter of interest can nonetheless be expressed as a counterfactual as the change in a descriptive parameter, the inequality in an adiposity measurement at age 10, that would be observed at the same age if we intervened on a variable which potentially mediates the relationship between ethnicity and the outcome (adiposity measurement). The change in inequality with and without the hypothetical intervention could be estimated by comparing the difference between the inequality when the intervention has been implemented and the original observed inequality without the intervention.

Despite the fact that variables such as diet, physical activity, and screen time are potential mediators of the relationship between ethnicity and adiposity measurements, it is more convenient to think of this estimate as the change in a descriptive quantity expected from a specified intervention. Sequential G-estimation can be used for this purpose because it allows for covariate adjustment of the intervention model while avoiding adjustment for the estimation of the inequality.25 This method first estimates a model of the effect of the intervention on the outcome using linear regression while adjusting for potential confounders which were selected using causal graphs and the relevant literature. In this study, confounders were collected through questionnaires (ethnic background, household income, maternal education, parity, marital status, maternal age, smoking and alcohol consumption during pregnancy, maternal prepregnancy body mass index, paternal body mass index), medical records (child sex, age, birth weight, gestational age, maternal weight gain during pregnancy), and physical measurements at the research center (child adiposity measurement at age 6). This model requires three causal assumptions: exchangeability (i.e., that there is no residual confounding or bias due to selection bias between

the proposed intervention and the measure of adiposity), positivity (i.e., that there is variation in the exposure conditional on confounders), and consistency (i.e., that the interventions is well defined). Consistency is likely not satisfied for screen time because we are not specifying an activity to replace screen time. However, this may be mitigated to an extent when we combine all interventions thereby replacing some screen time with sports participation. Using this model for the intervention, it is possible to estimate the value of the outcome if the intervention was set to different levels. Using this value of the outcome where the intervention is set to a specific level, the inequality had the intervention been implemented can be estimated. The difference between the inequality with and without the intervention is the change in ethnic inequality that would result from the intervention in question. Dividing this change by the inequality without the intervention gives the percent change in inequality. The results were expressed through the percent change in ethnic inequality and 95% confidence intervals (CIs). The code for the analysis, and which is designed to be easily repurposed by the user, can be found in the eAppendix; http://links.lww.com/EDE/B679.

We set diet score, screen time, and sports participation to the specific levels in each hypothetical intervention (Table 1). For the hypothetical extreme interventions, we used the maximum possible value and we decided on hypothetical cutoffs for the moderate interventions based on a combination of guidelines, categories available in our questionnaires and the distribution of the data.

For example, for the intervention setting diet score to five or greater, all children whose observed diet score was less than five were set to be equal to five. Equivalently, for the intervention where screen time is set to less than or equal to 2 hours/day, all children whose observed screen time was greater than 2 were set to be equal to 2. For the intervention setting sports participation to more than 4 hours a week, all children were set to the 4 or more hours a week category. An interaction term between ethnicity and each intervention was considered and retained if it had an important effect on the change in inequality. We also checked for interactions between the proposed interventions and sex.

We excluded twin births from the analysis (n = 243)to avoid data clustering and due to birth weight and developmental differences. We quantified losses to follow-up and contrasted the characteristics of the group that remained in

TABLE 1. Hypothetical Interventions Used

| Level | | Hypothetical Interventions | | | |
|----------|-------------------------|-----------------------------------|----------------------|----------------------------------|--|
| | Diet Score ^a | Screen Time | Sports Participation | Combined | |
| Extreme | =10 points | =0 hours/day | >4 hours/week | All three extreme interventions | |
| Moderate | ≥5 points | ≤2 hours/day | >2 hours/week | All three moderate interventions | |

the study with the group that was lost to verify whether the losses to follow-up were related to ethnic group, intervention, or outcome, which could introduce a selection bias.²⁶ To avoid selection bias associated with missing data, we performed multiple imputation by chained equations in R (15 imputed data sets, 1,000 iterations) using the fully conditional specification method.²⁷ More information can be found in eTable 1; http://links.lww.com/EDE/B679. Statistical analyses were performed using R version 3.6.1.

RESULTS

Sociodemographic characteristics at baseline were different between followed children and those lost to follow-up. For example, children who were lost to follow-up had lower birth weight and came from families with a non-Western ethnic background that had lower income and education (Table 2).

General characteristics of the sample after multiple imputation are shown in Table 3 (n = 9,506). Children from a Western ethnic background had a birth weight 130.6 g higher than children from a non-Western ethnic background. Mothers from a Western ethnic background were 2.6 years older and had a prepregnancy body mass index 1.6 kg/m² lower than mothers from a non-Western ethnic background, 12% more were nulliparous, and 15% more were married or living with a couple. Fathers from a Western ethnic background had a body mass index 0.5 kg/m² lower than fathers from a non-Western ethnic background. Both household income and maternal education were higher in families from Western than from non-Western ethnic backgrounds.

Table 4 shows that diet score, screen time, and sports participation were healthier in children from Western compared with non-Western ethnic backgrounds. All the adiposity measurements were lower in children from Western compared with

non-Western ethnic backgrounds at both 6 and 10 years old, and the difference was larger at age 10. The observed ethnic inequality at age 10 was -1.2 kg (95% CI = -1.7, -0.8) for weight, -1.1 kg/m^2 (CI = -1.2, -0.9) for body mass index (BMI), and -0.8 kg/m^2 (CI = -0.9, -0.7) for fat mass index (Table 5).

Table 5 also shows how different hypothetical lifestyle interventions would reduce ethnic inequalities in child adiposity at 10 years old. The interactions between sex and each intervention were all above P > 0.2 except for the screen time model with fat-free mass index as the outcome (P = 0.18). Based on these results, we chose not to stratify by sex. Reducing screen time was the specific intervention that mostly decreased ethnic inequalities in each measure of adiposity, followed by sports participation which also reduced inequalities but less so. Interventions related to diet score improvement did not appear to reduce ethnic inequalities in child adiposity. There was little difference in the amount of ethnic inequality reduced by the extreme and moderate interventions. The inequality in weight was the most reduced by the proposed interventions, followed by body mass index and fat mass index. There was little to no change in fat-free mass index.

Generally, combined diet, screen time, and sports participation interventions showed the largest reductions in ethnic inequalities in child adiposity. The most extreme intervention that would maximize the diet score, reduce screen time to 0 hours a day, and increase sports participation to over 4 hours a week showed reductions in ethnic inequalities by 21% (CI = 8%, 35%) for weight, 9% (CI = 4%, 14%) for body mass index, and 9% (CI = 6%, 13%) for fat mass index. Having a diet score of five points or higher, no more than 2 hours a day of screen time, and more than 2 hours a week of sports participation combined also reduced ethnic inequalities by 17% (CI = 6%, 28%) for weight, 7% (CI = 3%, 11%)

TABLE 2. Sociodemographic Characteristics of the Participants at Baseline, by Follow-up Status at Age 10

| | Followed at Age 10 | | Lost to Follow-up | |
|--|--------------------|-----------------|-------------------|-----------------|
| Characteristics | n | Mean (SD) | n | Mean (SD) |
| Birth weight (g) | 5,696 | 3,433.1 (553.6) | 3,727 | 3,373.5 (577.0) |
| Gestational age (weeks) | 5,669 | 39.9 (1.8) | 3,749 | 39.7 (2.1) |
| Maternal age (y) | 5,705 | 30.9 (5.0) | 3,799 | 28.3 (5.6) |
| Maternal weight gain during pregnancy (g/week) | 4,289 | 189.5 (238.3) | 2,644 | 193.8 (275.5) |
| Maternal prepregnancy body mass index (kg/m²) | 4,298 | 23.5 (4.1) | 2,655 | 23.9 (4.7) |
| Paternal body mass index (kg/m²) | 4,031 | 25.3 (3.4) | 2,147 | 25.4 (3.8) |
| | | n (%) | | n (% |
| Child sex (male) | 5,7 | 705 (50) | 3,7 | 98 (52) |
| Child ethnicity (Western ethnic background) | 5,561 (69) | | 3,306 (53) | |
| Household income (>1,600€ per month) | 4,418 (77) | | 2,192 (55) | |
| Maternal education (University) | 5,237 (50) | | 3,095 (30) | |
| Marital status (married/living together) | 5,234 (89) | | 3,109 (80) | |
| Parity (nulliparous) | 5,521 (57) | | 3,625 (52) | |
| Smoking during pregnancy (yes) | 4,979 (24) | | 3,045 (32) | |
| Alcohol consumption during pregnancy (yes) | 4,570 (57) | | 2,864 (40) | |

TABLE 3. Sociodemographic Characteristics at Baseline by Ethnicity After Multiple Imputation

| | Total $(n = 9,506)$ | Western Ethnic Background (n = 5,876) | Non-Western Ethnic Background (n = 3,630) | |
|--|---------------------|--|--|--|
| Characteristics | Mean (SD) | Mean (SD) | Mean (SD) | |
| Birth weight (g) | 3,404.4 (572.2) | 3,453.6 (569.0) | 3,324.8 (568.5) | |
| Gestational age (weeks) | 39.8 (1.9) | 39.9 (1.9) | 39.7 (2.0) | |
| Maternal age (y) | 29.9 (5.4) | 30.8 (5.0) | 28.3 (5.6) | |
| Maternal weight gain during pregnancy (g/week) | 191.6 (254.3) | 198.5 (236.6) | 180.3 (280.2) | |
| Maternal prepregnancy body mass index (kg/m²) | 23.7 (4.4) | 23.1 (3.9) | 24.7 (5.0) | |
| Paternal body mass index (kg/m²) | 25.4 (3.6) | 25.2 (3.4) | 25.7 (3.8) | |
| | n (%) | n (%) | n (%) | |
| Child sex (male) | 4,813 (51) | 2,966 (50) | 1,846 (51) | |
| Household income (>1,600€ per month) | 6,084 (64) | 4,795 (82) | 1,289 (35) | |
| Maternal education (University) | 3,869 (41) | 3,177 (54) | 692 (19) | |
| Marital status (married/living together) | 8,090 (85) | 5,322 (91) | 2,768 (76) | |
| Parity (nulliparous) | 5,230 (55) | 3,505 (60) | 1,725 (48) | |
| Smoking during pregnancy (yes) | 2,523 (27) | 1,577 (27) | 946 (26) | |
| Alcohol consumption during pregnancy (yes) | 4,749 (50) | 3,726 (63) | 1,022 (28) | |

TABLE 4. Diet, Screen Time, Sports Participation, and Adiposity Measurements in Children

| | Total $(n = 9,506)$ | Western Ethnic Background (n = 5,876) | Non-Western Ethnic Background (n = 3,630) | |
|--|---------------------|--|--|--|
| Characteristics | Mean (SD) | Mean (SD) | Mean (SD) | |
| Diet score at age 8 (points) | 4.4 (1.3) | 4.5 (1.2) | 4.3 (1.3) | |
| Screen time at age 9 (hours/day) | 2.8 (2.0) | 2.4 (1.6) | 3.5 (2.3) | |
| Weight at age 6 (kg) | 23.0 (4.8) | 22.6 (4.2) | 23.5 (5.7) | |
| Weight at age 10 (kg) | 34.9 (9.1) | 34.1 (7.7) | 36.0 (10.9) | |
| Body mass index at age 6 (kg/m ²) | 16.3 (2.2) | 16.0 (1.9) | 16.7 (2.5) | |
| Body mass index at age 10 (kg/m ²) | 17.6 (3.3) | 17.1 (2.8) | 18.3 (3.9) | |
| Fat mass index at age 6 (kg/m ²) | 4.1 (1.6) | 3.9 (1.4) | 4.4 (1.9) | |
| Fat mass index at age 10 (kg/m ²) | 4.8 (2.4) | 4.5 (2.0) | 5.4 (2.8) | |
| Fat-free mass index at age 6 (kg/m ²) | 12.0 (1.2) | 12.0 (1.2) | 12.1 (1.3) | |
| Fat-free mass index at age 10 (kg/m ²) | 12.6 (1.6) | 12.5 (1.4) | 12.8 (1.7) | |
| Sports participation at age 9 | % | % | % | |
| <1 hour/week | 6 | 4 | 9 | |
| 1–2 hours/week | 30 | 28 | 34 | |
| 2–4 hours/week | 43 | 45 | 40 | |
| >4 hours/week | 21 | 23 | 17 | |

for body mass index, and 7% (CI = 4%, 10%) for fat mass index. The combined diet, screen time, and sports participation interventions were not associated with changes in ethnic inequalities in fat-free mass index. Similar results were found when ethnicity was dichotomized into Dutch and non-Dutch ethnic background available in eTable 2; http://links.lww.com/EDE/B679.

DISCUSSION

We estimated that implementing combined diet, screen time, and sports participation interventions was the best option to reduce ethnic inequalities in child adiposity. Although ethnic inequality reductions were larger under the ideal intervention of a maximum diet score, 0 hours screen time, and more than 4 hours a week of sports participation, the strict nature of this type of intervention would be difficult to maintain over time. ^{28,29} The results of this study suggest that a less extreme intervention of a diet score of five points or higher, a screen time of 2 hours a day or less, and a sports participation of more than 2 hours a week reduced ethnic inequalities in weight, body mass index, and fat mass index in children at 10 years old by nearly the same amount.

TABLE 5. Reductions in Ethnic (Western/Non-Western) Inequalities on Adiposity from Hypothetical Diet, Screen Time, and Sports Participation Interventions in Children at Age 10

| | Ethnic Inequality (1: Western, 0: Non-Western) | | | | |
|---|--|----------------------------|---------------------------|--------------------------------|--|
| Interventions | Weight (kg) | Body Mass Index (kg/m²) | Fat Mass Index (kg/m²) | Fat-free Mass Index (kg/m²) | |
| Inequality without intervention | -1.21 (-1.65, -0.77) | -1.04 (-1.21,-0.88) | -0.79 (-0.91, -0.68) | -0.22 (-0.29, -0.14) | |
| Diet score = 10 | | | | | |
| Percentage change in inequality | -1% (-3%, 1%) | -1% (-1%, 0%) | 0% (-1%, 0%) | 0% (-2%, 3%) | |
| Inequality with intervention | $-1.20 \; (-1.64, -0.76)$ | -1.04 (-1.21, -0.87) | -0.79 (-0.90, -0.68) | -0.22 (-0.29, -0.14) | |
| Diet score ≥ 5 | | | | | |
| Percentage change in inequality | -1% (-2%, 1%) | 0% (-1%, 0%) | 0% (-1%, 0%) | 0% (-2%, 2%) | |
| Inequality with intervention | $-1.20 \; (-1.64, -0.76)$ | -1.04 (-1.21, -0.87) | -0.79 (-0.91, -0.68) | -0.22 (-0.29, -0.14) | |
| Screen time = 0 | | | | | |
| Percentage change in inequality | -16% (-27%, -5%) | -7% (-11%, -2%) | -7% (-10%, -3%) | -4% (-13%, 5%) | |
| Inequality with intervention | -1.02 (-1.46, -0.59) | -0.98 (-1.15, -0.81) | -0.74 (-0.85, -0.63) | -0.21 (-0.28, -0.13) | |
| Screen time ≤ 2 hours/day | | | | | |
| Percentage change in inequality | -13% (-23%, -4%) | -6% (-9%, -2%) | -6% (-8%, -3%) | -3% (-10%, 4%) | |
| Inequality with intervention | -1.05 (-1.49, -0.61) | -0.99 (-1.15, -0.82) | -0.75 (-0.86, -0.64) | -0.21 (-0.28, -0.14) | |
| Sports participation > 4 hours/week | | | | | |
| Percentage change in inequality | -6% (-10%, -2%) | -3% (-4%, -1%) | -3% (-4%, -1%) | 0% (-2%, 2%) | |
| Inequality with intervention | -1.15 (-1.59, -0.70) | -1.02 (-1.19, -0.85) | -0.77 (-0.89, -0.66) | -0.22 (-0.29, -0.14) | |
| Sports participation > 2 hours/week | | | | | |
| Percentage change in inequality | -4% (-7%, -1%) | -2% (-3%, -1%) | -2% $(-3%, -1%)$ | 0% (-2%, 2%) | |
| Inequality with intervention | -1.16 (-1.61, -0.72) | -1.03 (-1.19, -0.86) | -0.78 (-0.89, -0.67) | -0.22 (-0.29, -0.14) | |
| Combined: Diet score = 10, screen time | = 0, and sports participation > | 4 hours/week | | | |
| Percentage change in inequality | -21% (-35%, -8%) | -9% (-14%, -4%) | -9% (-13%, -6%) | -4% (-13%, 6%) | |
| Inequality with intervention | $-0.96 \; (-1.41, -0.52)$ | -0.95 (-1.12, -0.78) | -0.72 (-0.84, -0.61) | -0.21 (-0.28, -0.13) | |
| Combined: Diet score \geq 5, screen time \leq | ≤ 2 hours/day, and sports partici | pation > 2 hours/week | | | |
| Percentage change in inequality | -17% (-28%, -6%) | -7% (-11%, -3%) | -7% (-10%, -4%) | -3% (-11%, 5%) | |
| Inequality with intervention | $-1.01 \; (-1.46, -0.57)$ | -0.97 (-1.14, -0.80) | -0.74 (-0.85, -0.62) | -0.21 (-0.29, -0.14) | |

Estimates expressed with 95% confidence intervals in parentheses.

The combined extreme intervention was able to reduce the inequality in weight by 17% and the inequality in body mass index by 8%. The reason these reductions are not larger is possibly because they are applied simultaneously to both children from Western and non-Western ethnic backgrounds. The degree to which the inequality is reduced is a function of how much more the children from a non-Western ethnic background benefit from the intervention relative to those from a Western ethnic background. The reduction would have been larger if it had been only targeted toward children from a non-Western ethnic background, but such a targeted intervention is a less realistic public health intervention than targeting all children. Also, the observed change in ethnic inequalities was over a 4-year period (between 6 and 10 years old). It may be that longer periods of time are required to make larger gains in reducing ethnic health inequalities.

Measurement error in the intervention variables may also bias the association between the intervention and the outcome toward the null³⁰ resulting in underestimation of the change in health inequality. Related to this is the fact that some

interventions were on categorical variables. Any residual differences within a category (e.g., if children from a Western ethnic background in the highest physical activity group did more exercise than children from a non-Western ethnic background in the highest physical activity group) would not be captured in the model and thus would not contribute to the reduction in inequalities.

Screen time was associated with the highest reduction in ethnic inequalities when applied independently. This is important because screen time has increased dramatically during the last decades, 31,32 and current evidence suggests that screen media exposure leads to obesity in children and adolescents mainly through increased eating while viewing, exposure to high-calorie, low-nutrient food and beverage marketing that influences children's preferences, purchase requests, and consumption habits.^{33,34} However, this intervention by itself is not completely well defined (i.e., does not satisfy the consistency assumption) because it depends on what a child will do with the time they would normally have been in front of a screen. The more specific the intervention the easier it is to generalize results and guide the elaboration and evaluation of public health interventions. In this study, combined interventions partially address this issue by specifying that while the child is reducing screen time they are also increasing sport participation and improving diet quality.

Currently, comprehensive evidence is lacking on what works in terms of interventions to reduce inequalities in childhood obesity.35-37 A systematic review that examined the best available international evidence from interventions to reduce socioeconomic inequalities in obesity-related outcomes among children found that individual-level, screen time reduction, and mentor-based health promotion interventions could be effective in reducing inequalities in childhood obesity because they were more beneficial for low socioeconomic children who already had obesity. For community- and societal-level interventions, the evidence was inconclusive.³⁸

Although body mass index is the most common measure of obesity at the population level it fails to distinguish between fat and fat-free mass and may exaggerate obesity in large muscular children.³ For this reason, we also included measurements such as fat mass index and fat-free mass index to have more complete information about the child's body composition. We observed that with the hypothetical interventions all measurements related with child adiposity had the same direction of effect, they consistently decreased in both ethnic groups. When looking at relative reductions, we observed that the decrease was higher in the group of children with a non-Western ethnic background than with a Western ethnic background, which translated into a reduction of the ethnic inequality. With respect to fat-free mass index, the change in ethnic inequality with the hypothetical interventions did not change much from the ethnic inequality without the interventions. This suggest that the hypothetical interventions were more associated with reducing measures of adiposity than fatfree mass.

In this study, there were differences by ethnicity in all adiposity measurement among children at birth, 6 years old, and 10 years old, and these differences increased over time, with less healthy levels for children from non-Western than Western ethnic backgrounds. A European systematic review of the literature concluded that children of migrants were more at risk of overweight and obesity than native children.⁴ Previous studies conducted in the Netherlands found that prevalence of overweight and obesity was especially alarming among children in the two residing largest migrant groups: those from Turkish and Moroccan descent. 39,40

There were also differences in diet, screen time, and sports participation by ethnicity, and children from a non-Western ethnic background had less healthy behaviors. It has been described that migrants tend to abandon their traditional food habits, adopt westernized dietary patterns, and engage in a more sedentary way of life. 41,42

Alterations in physical activity, leading to a more sedentary way of life, have been reported in immigrants compared

with women and men born in Sweden.⁴³ Also, migrant children in European countries displayed lower levels of physical activity than native ones. 4,44,45 A systematic review about physical activity among North African immigrant children in European countries found that one of the major predisposing factors to obesity is the lack of a health-conscious exercise culture. 46 Another aspect responsible for sedentary lifestyle during childhood is the high exposure to media, in particular television, which exposes children to advertising for highenergy snacks and reduces physical activity.⁴¹ The evidence suggests that daily screen time for more than 2 hours is associated with unfavorable body composition and that lowering sedentary time leads to reductions in body mass index and health risk both in children and youth.⁴⁷

With respect to dietary habits, a systematic review showed that some ethnic groups in Europe are likely to become less healthy as individuals increase consumption of processed foods that are energy dense and contain high levels of fat, sugar, and salt. After migration, the majority of ethnic groups observed alter their eating habits, combining parts of their traditional diet with some of the less healthy elements of the Western diet, such as soft drinks, sweets, and snacks that are higher in fat or sugar and lower in dietary fiber than the legumes and grains of the traditional diet. These mixed food habits are emerging mainly among younger people, probably as the consequence of acculturation processes, which increases the risk of overweight and obesity.⁴¹

Ethnic inequalities in child adiposity might be explained at individual level by the lack of exercise, increasing consumption of less healthy foods, genetic predisposition, and stress related to migration and settlement of the family. Additionally, acculturation and stronger integration within the new host country have been associated with a decline in the health of immigrants. 41,46 Other factors are income, food availability, religion, food beliefs, body size preferences, age, immigrant generation, and country of origin. 41,48-51 Therefore, effective public health interventions aimed to reduce ethnic inequalities in child adiposity need to be culturally sensitive and these only can be adequately designed, if particular lifestyle behaviors that influence the elevated risk among specific ethnic groups are identified.48

This is one of the first studies using this methodology to quantify the reductions in ethnic inequalities in child adiposity at 10 years old from hypothetical diet, screen time, and sports participation interventions. There have been important advances in methodology in the study of health inequalities in the context of causal inference recently^{23,24} that have helped researchers target the parameter of interest: how interventions decrease inequalities. Whereas many studies adjust the ethnic inequality for confounders of the ethnicity-outcome relationship, ^{24,25} this study employs special methods to avoid this. The rationale is that health inequalities are of interest unadjusted for other variables because the interest is in real-world differences in health variables by ethnicity, not the difference after adjusting for confounders. Therefore, in analyses that estimate the reduction in a health inequality due to a specific intervention, the interest is in the reduction in health disparity unadjusted for covariates. By employing these methods, the results of this research help to better understand how interventions could reduce ethnic inequalities in children and to be realistic about what effect size to expect. Other strengths include that the results obtained came from a population-based, multicultural cohort study and that multiple imputation was performed to avoid potential selection bias and to have representative information.

Among the limitations of this study is that information recall and social desirability bias could affect the food frequency and screen time questionnaires. 52,53 Recall bias may occur because the food frequency questionnaire required participants to remember their diet during the last 4 weeks, which cannot be done with perfect accuracy, and the social desirability bias refers to the tendency to give response that are viewed more favorably by the interviewer. Another potential limitation is the cross-cultural validity of the diet score, because for religious or other beliefs, specific ethnic groups may avoid certain foods that are considered healthier and this could make one group's diet score appear better than another group's.⁵⁴ Also, the presence of differential missing data by ethnicity may lead to an underestimation of the ethnic inequalities in child adiposity. We used multiple imputation to ensure that all participants were included in the analysis to avoid this bias. Nonetheless, the results could still be biased to the degree that missingness is related to variables not used in the imputation. We attempted to minimize this by including all variables from the analyses in the imputation and additional variables found to be related to missingness. Last, unmeasured confounding of the intervention-outcome model would lead to an overestimate of the change in ethnic inequalities in child adiposity.

This study generates evidence that could help, in conjunction with other research, to guide public health interventions to reduce ethnic inequalities in child health. For future research, it would be interesting to use this methodology to evaluate the effect of public health interventions on cardiometabolic risk. Also, longitudinal models could be developed to follow participants over time to identify the age at which interventions to reduce ethnic inequalities are most effective.

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