Modelling Outcomes of Complex Treatment Strategies Following a Clinical Guideline for Treatment Decisions in Patients with Rheumatoid Arthritis

An Tran-Duy · A. M. M. van Riel · Mart A. F. J. van de Laar · Johan L. Severens

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Abstract Background Management of rheumatoid arthritis (RA) is characterised by a sequence of disease-modifying anti-rheumatic drugs (DMARDs) and biological response modifiers (BRMs). In most of the Western countries, the drug sequences are determined based on disease activity and treatment history of the patients. A model for realistic patient outcomes should reflect the treatment pathways relevant for patients with specific characteristics. Objective This study aimed at developing a model that could simulate long-term patient outcomes and cost effectiveness of treatment strategies with and without inclusion of BRMs following a clinical guideline for treatment decisions. Methods Discrete event simulation taking into account patient characteristics and treatment history was used for model development. Treatment effect on disease activity, costs, health utilities and times to events were estimated using Dutch observational studies. Long-term progression of physical functioning was quantified using a linear mixed-effects model. Costs and health utilities were estimated using two-part models. The treatment strategy recommended by the Dutch Society for Rheumatology where both DMARDs and BRMs were available (Strategy 2) was compared with the treatment strategy without BRMs (Strategy 1). Ten thousand theoretical patients were tracked individually until death. In the probabilistic sensitivity analysis, Monte Carlo simulations were performed with 1,000 sets of parameters sampled from appropriate probability distributions.

Results The simulated changes over time in disease activity and physical functioning were plausible. The incremental cost per quality-adjusted life-year gained of Strategy 2 compared with Strategy 1 was €124,011. At a willingness-to-pay threshold higher than €119,167, Strategy 2 dominated Strategy 1 in terms of cost effectiveness but the probability that the Strategy 2 is cost effective never exceeded 0.87.

Conclusions It is possible to model the outcomes of complex treatment strategies based on a clinical guideline for the management of RA. Following the Dutch guideline and using real-life data, inclusion of BRMs in the treatment

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strategy for RA appeared to be less favourable in our model than in most of the existing models that compared drug sequences independent of patient characteristics and used data from randomised controlled clinical trials. Despite complexity and demand for extensive data, our modelling approach can help to identify the knowledge gaps in clinical guidelines for RA management and priorities for future research.

**Key Points for Decision Makers**

Treatment strategies for patients with early-stage rheumatoid arthritis (RA) are becoming increasingly complex and it is therefore necessary to develop a modelling approach that fully supports economic evaluation of a wide range of real-life treatment strategies.

Treatment strategies including biological response modifiers (BRMs) with treatment decisions based on patient characteristics and treatment history have been accepted in most of the Western countries, but few cost-effectiveness models have attempted to take the use of BRMs in clinical practice into consideration.

Modelling the outcomes of complex treatment strategies based on a clinical guideline for the management of RA is possible and helpful in assessing the impact of changes in clinical practice on disease progression, health and costs of the RA patients, and cost effectiveness of the new treatment decisions.

Following the Dutch guideline for management of RA and using real-life data, the simulated incremental cost per quality-adjusted life-year gained of the treatment strategy including BRMs compared with the treatment strategy without BRMs was higher than conventionally accepted willingness-to-pay thresholds, and the probability that the treatment strategy including BRMs is cost effective never exceeded 0.87.

**1 Introduction**

In the management of rheumatoid arthritis (RA), early treatment with synthetic disease-modifying antirheumatic drugs (DMARDs) and subsequent use of biological response modifiers (BRMs) to achieve low disease activity or remission have been recommended for clinical practice in both American and European treatment guidelines [1-4]. Because BRMs are far more expensive than DMARDs, the incremental costs per quality-adjusted life-year (QALY) [5] gained (ICERs) of treatments with BRMs have been discussed extensively.

Despite the growing number of published studies on the cost effectiveness of BRMs, few have attempted to take the use of BRMs in clinical practice into consideration. The sequence of drugs in the real-life management of RA is determined based on characteristics, notably disease activity using the Disease Activity Score for 28 joints (DAS28) [3, 6, 7], and treatment history of the patients. It is well recognised that the effectiveness of different drugs and of the same drug at different positions in a drug sequence are most likely different. Additionally, early or late administration of a drug affects the course of disease and thus the health and costs of the patients. Therefore, a model for realistic patient outcomes should reflect the treatment pathways relevant for patients with specific characteristics. Recently, there has been an increasing tendency of adopting sequential treatment strategies instead of single drugs as comparators in the cost-effectiveness models for RA in different European countries [8-15]. However, the order of the drugs in the existing models was determined independently of disease activity and treatment history of the patients.

The present study aimed at developing a model that could simulate long term patient outcomes and cost effectiveness of treatment strategies with and without inclusion of BRMs; following a formal clinical guideline and rheumatologist opinion for treatment decisions based on patient characteristics and treatment history, and using real-life data to parameterize the model.

**2 Methods**

2.1 Model Structure

Discrete event simulation (DES) was used to track changes over time in disease measures, utilities and costs of the individual patients. This modelling approach can overcome intractable problems in conventional models for a chronic disease (e.g. numerous chance nodes in a decision tree or a large number of states in a Markov model) [16-18]. Moreover, memorising changes in patient characteristics over time is possible in DES, which permits simulation of the rheumatologist decisions on starting and switching treatments based on patient history [19].

The main components of the DES model consist of entities, states and events. An entity is an RA patient with a set of characteristics (e.g. age, gender, rheumatic factor, disease duration, disease activity and physical functioning), which influence the simulation outcomes.
States in our model were defined based on trends of change in DAS28 and by on-going treatments to help determine potential events that may occur during tracking disease progression and treatment decisions. Different from those in a Markov model, these states are not necessarily mutually exclusive. Conceptually, a patient can be in three different disease activity phases (DAS28-related states) while treated with a specific drug [20]. First, the decreasing phase is characterized by a steady decrease in DAS28 as a response to the treatment. Second, the maintenance phase is characterized by small fluctuations of DAS28 surrounding a constant level. Third, the increasing phase is characterized by a continuous increase in DAS28 after the patient stops responding to the drug. In this case, DAS28 was assumed to return to the patient’s baseline level within 12 weeks. When the patient is receiving ‘palliative’ therapy, i.e. the use of only low-dose corticosteroids after failure of all drugs in a treatment strategy, DAS28 was assumed to be stable. These DAS28-related states were used to determine the events that may occur given a trend of change in DAS28 (see Table 1).

Regarding the treatments, a patient can be in one of eight states, referred to as treatment-related states, which is being treated with the first or second DMARD, the first or second tumour necrosis factor α (TNF) inhibitor BRM, the first or second non-TNF BRM, the DMARDs after failure of the first two DMARDs, or ‘palliative’ therapy. Treatment-related states were used to determine changes in DAS28, times to DAS28-related events and a new treatment when the current drug fails.

Events in the model were classified as DAS28-related events and DAS28-neutral events. DAS28-related events consisted of “End of DAS28 decrease”, “Loss of response to the current drug” (which caused an increase in DAS28) and “DAS28 reaching 1.2 unit higher”. The last event was formulated to help calculate the rate of DAS28 increase based on the assumption that DAS28 returned to the baseline level in 12 weeks after a loss of response to the current treatment, and that a 1.2-unit change in DAS28 was significant [21]. DAS28-neutral events consisted of “Severe toxicity of the current drug”, “Visit a rheumatologist”, “Select a new treatment”, “Start a new treatment” (i.e. first administration of a new treatment) and “Death”. A summary of the interdependence among states and events is given in Table 1. For competing events, the patient will “jump” to the event to which the sampled time is shortest. When an event occurred, an associated procedure was invoked for implementation where the patient characteristics were updated and times to the next events were computed. It should be noted that “Select a new treatment” and “Start a new treatment” may immediately follow “Visit a rheumatologist”, but “Visit to a rheumatologist” does not necessarily lead to the occurrence of “Select a new treatment” and therefore not “Start of a new treatment” as the next events. This is the reason for a separation of these three events. The general simulation process is depicted in Fig. 1. Details on the simulation process can be found in the online supplementary appendix.

2.2 Treatment Strategies

We compared the treatment strategy recommended by the Dutch Society for Rheumatology where both DMARDs and BRMs were available (Strategy 2) with the treatment strategy without inclusion of BRMs (Strategy 1). Availability and order of the drugs in these two strategies were as follows:

- Strategy 1: eight DMARDs available in the following sequence: methotrexate followed randomly by sulphasalazine or leflunomide, which was followed by azathioprine, cyclosporin A, cyclophosphamide, hydroxychloroquine and injectable gold at a random order;

<table>
<thead>
<tr>
<th>Table 1 Interdependency of the model states and events</th>
</tr>
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<tbody>
<tr>
<td>DAS28-related state</td>
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<tr>
<td>--------------------</td>
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<tr>
<td>On DAS28 decrease</td>
</tr>
<tr>
<td>On DAS28 maintenance</td>
</tr>
<tr>
<td>On DAS28 increase</td>
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<tr>
<td>On DAS28 stability</td>
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DAS28 Disease Activity Score for 28 joints, DMARD disease-modifying antirheumatic drug, TNF tumour necrosis factor, BRMs biological response modifiers, ‘palliative’ therapy the use of only low-dose corticosteroids after failure of all drugs in a treatment strategy

<sup>a</sup> Competing events

<sup>b</sup> Events that occur only when a visit to a rheumatologist occurs

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Fig. 1 Flow chart of the general simulation process. The triangle at the top represents the entry of the simulation. The rectangles with double-line and single-line borders represent procedures and instructions, respectively. The diamonds represent decisions. The boxes with dotted borders provide extra information for the flow chart. See text and the online supplementary appendix for details. DAS28 Disease Activity Score for 28 joints, HAQ Health Assessment Disability Index

- Strategy 2: eight DMARDs as those in Strategy 1, and two anti-TNF and two non-TNF BRMs, available in the following sequence: methotrexate followed randomly by sulphasalazine or leflunomide, two anti-TNF BRMs, and two non-TNF BRMs, which were followed by azathioprine, cyclosporin A, cyclophosphamide,
hydroxychloroquine and injectable gold at a random order. The two anti-TNF BRMs comprised one monoclonal antibody and one soluble receptor fusion protein, which were randomly chosen from etanercept, adalimumab, infliximab, golimumab and certolizumab. The two non-TNF BRMs were randomly chosen from rituximab, abatacept and tocilizumab. For patients with a positive rheumatic factor, rituximab must be available at the first position of the non-TNF BRMs; for patients with a negative rheumatic factor, abatacept or tocilizumab must be available at the first position of the non-TNF BRMs.

The order of the drugs and random selections between sulphasalazine and leflunomide, among anti-TNF BRMs, among non-TNF BRMs and among DMARDs other than sulphasalazine and leflunomide in the above-mentioned treatment strategies agreed well with those recorded in the observational cohorts used for model inputs (see Sect. 2.3). The algorithm for treatment decisions in the model was formulated based on the Dutch guidelines and rheumatologist opinion for the treatment of RA. First, a flow chart of treatment decisions was drawn based on the general guidelines and delivered to the rheumatologists at the Universities of Maastricht, Nijmegen and Twente in The Netherlands. Second, several meetings were organized among the stakeholders to discuss and refine the flow chart based on personal experience and knowledge about daily clinical practice. Finally, the algorithm was concluded based on a consensus on the treatment decisions that reflect the predominant clinical practice of the management of RA in The Netherlands. In this algorithm, methotrexate, in combination with intramuscularly injected corticosteroids to bridge the time of action on disease activity of the first DMARD, was started as soon as RA is diagnosed. The next treatment based on the order of the drugs in the respective strategy was considered when a drug failed primarily or secondarily, or caused severe toxicity. Primary failure was assumed if DAS28 was still higher than 3.2 after 3–6 months (random) since the start of the treatment. Secondary failure was assumed when DAS28 went back to a level >3.2 after a primary response to the treatment. In Strategy 2, BRMs were combined with methotrexate if the patient did not experience severe toxicity when receiving methotrexate monotherapy after RA diagnosis. When there were no DMARDs left in the sequence, 'palliative' therapy was given in both strategies.

2.3 Model Inputs and Data Sources

2.3.1 Patients’ Baseline Characteristics

The initial cohort to be simulated consisted of patients who were newly diagnosed with RA. We used the registry of the Radboud University Nijmegen Medical Centre (Nijmegen Inception Cohort) [22] to estimate the empirical distributions of baseline age and gender, DAS28 and the presence of positive rheumatic factor. This observational study was started in 1985 to collect information on the long-term course of RA in patients with disease duration <1 year and being DMARD naïve on the inclusion date. Because the course of disease in RA patients has become milder in recent years [23], only data of patients who entered the registry after 2002 were used for parameter estimation. Between January 1st, 2003 and May 22nd, 2012, there were 231 patients with a mean follow-up of 1.5 years (minimum 0 year; maximum 9.2 years).

Percentages of male and female patients with paid jobs were estimated based on a Dutch cross-sectional study on labour force participation [24]. Based on these distributions, the baseline characteristics of each patient were sampled using the sampling methods described in the online supplementary appendix.

2.3.2 Treatment Effectiveness and Toxicity

Data from the Dutch Rheumatoid Arthritis Monitoring (DREAM) biologic registry [25] (BRM treatment) and the Nijmegen Inception cohort (DMARD treatment) were used to estimate: treatment effect on DAS28, time to end of DAS28 decrease and time to occurrence of severe toxicity since the start of a drug; and time to loss of response since the end of DAS28 decrease. The DREAM biologic registry was started in April 2003 to monitor and evaluate the use of BRMs in patients who had not responded to methotrexate (optimal dose) and at least another DMARD before the inclusion date from 13 hospitals in The Netherlands. Up to June 28th, 2011, 1,799 patients had been included in this cohort with a mean follow-up of 2.2 years (minimum 0 year; maximum 8.2 years).

Because of insufficient data, we assumed that the effectiveness of a specific drug was independent of the identities and the causes of failure of the drugs that had been given previously. We found a linear relationship between the absolute change in DAS28 during the decreasing phase and DAS28 at the start of a treatment. For the same BRM, we found different effectiveness between its first and second administration. Therefore, absolute changes in DAS28 were sampled for each drug or drug class, distinguishing the first and second BRM, using a statistical linear model with DAS28 at the start of the treatment as an explanatory variable (Table 2). The assumption on normal distribution of the errors was verified using the Shapiro–Wilks test [26]. Because this assumption was satisfied for any linear model for DAS28 change, the assumption on homoscedasticity of the errors was further verified using the Breusch–Pagan test [27].
Table 2 Model parameters for sampling patient characteristics and times to events, and estimating treatment effectiveness, and costs and QALYs

<table>
<thead>
<tr>
<th>Parameters</th>
<th>Estimate</th>
<th>Distribution*</th>
<th>Source</th>
</tr>
</thead>
<tbody>
<tr>
<td>Proportion of male patients who have paid jobs</td>
<td>0.85</td>
<td>Beta(249, 43)</td>
<td>Chorus et al. [24]</td>
</tr>
<tr>
<td>Proportion of female patients who have paid jobs</td>
<td>0.42</td>
<td>Beta(318, 444)</td>
<td>Chorus et al. [24]</td>
</tr>
<tr>
<td>Intercept (α) and slope (β) in the linear regression model for estimating absolute change in DAS28 based on DAS28 at the start of:</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Methotrexate (n = 91)</td>
<td>α = −1.08; β = 0.73</td>
<td>MVN</td>
<td>NIC</td>
</tr>
<tr>
<td>Sulphasalazine (n = 27)</td>
<td>α = −0.98; β = 0.68</td>
<td>MVN</td>
<td>NIC</td>
</tr>
<tr>
<td>Leflunomide (n = 21)</td>
<td>α = −1.42; β = 0.71</td>
<td>MVN</td>
<td>NIC</td>
</tr>
<tr>
<td>Other DMARDs (n = 18)</td>
<td>α = −0.67; β = 0.32</td>
<td>MVN</td>
<td>NIC</td>
</tr>
<tr>
<td>Adalimumab, first position in sequential use of anti-TNF BRMs (n = 226)</td>
<td>α = −0.95; β = 0.72</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Adalimumab, second position in sequential use of anti-TNF BRMs (n = 120)</td>
<td>α = −1.48; β = 0.75</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Etanercept or certolizumab, first position in sequential use of anti-TNF BRMs (n = 267)</td>
<td>α = −0.81; β = 0.70</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Etanercept or certolizumab, second position in sequential use of anti-TNF BRMs (n = 167)</td>
<td>α = −0.81; β = 0.53</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Infliximab or golimumab, first position in sequential use of anti-TNF BRMs (n = 188)</td>
<td>α = −0.33; β = 0.46</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Infliximab or golimumab, second position in sequential use of anti-TNF BRMs (n = 19)</td>
<td>α = −0.36; β = 0.41</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Non-TNF BRM, first position in sequential use of non-TNF BRMs (n = 73)</td>
<td>α = −0.71; β = 0.57</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Non-TNF BRM, second position in sequential use of non-TNF BRMs (n = 18)</td>
<td>α = −0.72; β = 0.53</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Coefficients in the linear mixed models for predicting HAQ, HAQ = (α + u_a) + β_1 \times DAS28 + β_2 \times AGE + (β_3 + u_i) \times DR + β_4 \times SEX + ε in patients receiving a:</td>
<td></td>
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</tr>
<tr>
<td>DMARD (n = 185)</td>
<td>α = −0.39; β_1 = 0.14; β_2 = 0.01; β_3 = 0.03; β_4 = 0.29</td>
<td>MVN</td>
<td>NIC</td>
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<tr>
<td>BRM (n = 1,579)</td>
<td>α = −0.31; β_1 = 0.15; β_2 = 0.01; β_3 = 0.01; β_4 = 0.24</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Scale parameter in the exponential distribution of time to end of DAS28 decline (in year) in patients receiving u/an:</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>DMARD (n = 157)</td>
<td>1.69</td>
<td>N(0.53, 0.09)</td>
<td>NIC</td>
</tr>
<tr>
<td>Anti-TNF BRM (n = 987)</td>
<td>1.61</td>
<td>N(0.48, 0.53)</td>
<td>DREAM</td>
</tr>
<tr>
<td>Non-TNF BRMs (n = 91)</td>
<td>0.49</td>
<td>N(−0.71, 0.170)</td>
<td>DREAM</td>
</tr>
<tr>
<td>Scale parameter in the exponential distribution of time to loss of response (in year) of patients receiving a:</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>DMARD (n = 121)</td>
<td>0.97</td>
<td>N(−0.03, 0.18)</td>
<td>NIC</td>
</tr>
<tr>
<td>Anti-TNF BRMs (n = 726)</td>
<td>2.38</td>
<td>N(0.87, 0.13)</td>
<td>DREAM</td>
</tr>
<tr>
<td>Non-TNF BRMs (n = 52)</td>
<td>0.58</td>
<td>N(−0.54, 0.17)</td>
<td>DREAM</td>
</tr>
<tr>
<td>Coefficients in the parametric survival regression model for time to severe toxicity (in year), log(π) = α + β \times AGE, of:</td>
<td></td>
<td></td>
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</tr>
<tr>
<td>Methotrexate (n = 139)</td>
<td>α = 2.46; β = −0.021</td>
<td>MVN</td>
<td>NIC</td>
</tr>
<tr>
<td>Sulphasalazine (n = 42)</td>
<td>α = 2.95; β = −0.025</td>
<td>MVN</td>
<td>NIC</td>
</tr>
<tr>
<td>Leflunomide (n = 39)</td>
<td>α = −0.32; β = −0.014</td>
<td>MVN</td>
<td>NIC</td>
</tr>
<tr>
<td>Other DMARDs (n = 23)</td>
<td>α = 2.44; β = −0.019</td>
<td>MVN</td>
<td>NIC</td>
</tr>
<tr>
<td>Anti-TNF BRMs, first position in sequential use of anti-TNF BRMs (n = 1,136)</td>
<td>α = −0.62; β = −0.012</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Anti-TNF BRMs, second position in sequential use of anti-TNF BRMs (n = 510)</td>
<td>α = 2.24; β = −0.026</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
</tbody>
</table>

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analysed all the DMARDs that a patient received after failure of methyltrexate and sulfasalazine (or leflunomide) as one class, referred to as 'other DMARDs', because the numbers of observations for some drugs were too small ($n < 5$) and the effectiveness of the drugs with more observations were found to be similar; the effect of 'other DMARDs' on DAS28 was assumed to be the same in both treatment strategies 1 and 2. For the same reasons, all non-TNF BRMs at the same position in the treatment sequence were grouped, and specific pairs of anti-TNFs also grouped (see Table 2).

Table 2 continued

<table>
<thead>
<tr>
<th>Parameters</th>
<th>Estimate</th>
<th>Distribution</th>
<th>Source</th>
</tr>
</thead>
<tbody>
<tr>
<td>Non-TNF BRMs, first position in sequential use of non-TNF BRMs ($n = 121$)</td>
<td>$\alpha = -0.21; \beta = 0.019$</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Non-TNF BRMs, second position in sequential use of non-TNF BRMs ($n = 32$)</td>
<td>$\alpha = -2.98; \beta = 0.072$</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Coefficients in the two-part model for predicting 3-month non-drug healthcare cost (€) ($n = 425$)</td>
<td></td>
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<tr>
<td>1st part: logit($P(\text{non-zero cost})$) = $\alpha + \beta_1 \times DR + \beta_2 \times HAQ + \beta_3 \times HAQ \times SEX + \beta_4 \times HAQ \times SEX$</td>
<td>$\alpha = -0.87; \beta_1 = 0.02; \beta_2 = 0.78; \beta_3 = -0.31; \beta_4 = 0.67$</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>2nd part: logit(cost) = $\alpha_1 + \beta_1 \times DAS28 + \beta_2 \times HAQ + \beta_3 + \beta_4 \times DAS28 \times SEX + \epsilon$</td>
<td>$\alpha_1 = 5.93; \beta_1 = 0.18; \beta_2 = 0.42; \beta_3 = 0.93; \beta_4 = -0.24$</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>Coefficients in the two-part model for predicting 3-month sick leave cost (€) ($n = 425$)</td>
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<tr>
<td>1st part: logit($P(\text{non-zero cost})$) = $\alpha + \beta_1 \times AGE + \beta_2 \times SEX$</td>
<td>$\alpha = 1.90; \beta_1 = -0.07; \beta_2 = -0.91$</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>2nd part: logit(cost) = $\alpha + \beta_1 \times HAQ + \beta_2 + \beta_3 \times SEX + \epsilon$</td>
<td>$\alpha = 7.87; \beta_1 = 0.85; \beta_2 = -1.53$</td>
<td>MVN</td>
<td>DREAM</td>
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<tr>
<td>Coefficients in the two-part model for predicting health utility (HU) ($n = 425$)</td>
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</tr>
<tr>
<td>1st part: logit($P(\text{class 2})$) = $\alpha + \beta_1 \times AGE + \beta_2 \times DAS28 + \beta_3 \times HAQ$</td>
<td>$\alpha = 4.33; \beta_1 = 0.04; \beta_2 = -0.56; \beta_3 = -1.69$</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>2nd part: $HU_{\text{class 1}} = \alpha + \beta_1 \times HAQ + \beta_2 \times SEX + \epsilon$</td>
<td>$\alpha = 0.17; \beta_1 = -0.11; \beta_2 = 0.04$</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
<tr>
<td>$HU_{\text{class 2}} = \alpha + \beta_1 \times HAQ + \beta_2 \times SEX + \beta_3 \times DAS28 + \epsilon$</td>
<td>$\alpha = 0.82; \beta_1 = -0.11; \beta_2 = 0.03; \beta_3 = -0.01$</td>
<td>MVN</td>
<td>DREAM</td>
</tr>
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</table>

**DMARD** disease-modifying antirheumatic drug; **TNF** tumour necrosis factor; **BRM** biological response modifier; **SEX** takes on a value of 1 or 0 if the patient is female or male, respectively; **DR** disease duration; **HAQ** Health Questionnaire Assessment Disability Index; **DAS28** Disease Activity Score for 28 joints; **n** number of observations in the data set used for model fitting; **NIC** Nijmegen Inception Cohort; **DREAM** Dutch Rheumatoid Arthritis Monitoring registry

- **Distribution** used to sample parameter values in the probabilistic sensitivity analysis; $M(n, z)$, normal distribution with mean $m$ and standard deviation $z$; Beta(a, b), beta distribution with scale parameters a and b, which were estimated using the method of moments [55]; MVN, multivariate normal distribution with mean and variance-covariance matrix estimated by fitting the regression models
- $n_0$ and $n_1$ followed bivariate normal distribution with a mean vector $0$ and the variance-covariance matrix estimated by fitting the linear mixed model
- **Distribution** of the logarithm of the scale parameter; see Selvin [57] for method of parameter estimation
- $\mu$ is the scale parameter in the exponential distribution of time to toxicity

Long-term progression of physical functioning, measured by the Health Assessment Questionnaire Disability Index (HAQ) [29–31], was simulated based on the longitudinal relationship between DAS28 and HAQ. The changes in DAS28 and HAQ over time were linked to costs and

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health utility (HU) using statistical models (see next sections and Table 2).

2.3.3 Costs and Quality-Adjusted Life-years

Data from the DREAM biologic registry were used for statistical modelling of healthcare and sick leave costs and HU. Costs incurred by each patient were calculated based on the amounts of resource use and unit costs. The amounts of resource use were reported every 3 months by each patient participating in the DREAM registry and included: numbers of diagnostic and laboratory tests; days in hospitals; numbers of surgical procedures; visits to rheumatologists and other specialists, to general practitioners, to nurse specialists, to physiotherapists, and to psychologists; and hours of formal and informal care. The unit costs published by the Dutch Health Care Insurance Board [32] were used and adjusted using a consumer price index [33]. The human capital approach was used for estimating sick leave cost [34].

For healthcare and sick leave costs, we used two-part models as recommended by Mullahy [35] and Manning and Mullahy [36] to link cumulative costs over 3 months to age, gender, disease duration and mean DAS28 and HAQ. Because of the excessive zeros in the cost data, logistic regression models were used in the first part to estimate the probability that the costs were non-zero. Conditional on non-zero costs, log-linear regression models were used in the second part to evaluate the costs. The assumptions on normality and homoscedasticity of the errors in these log-linear models were verified using the same tests as in the linear models for changes in DAS28. Because the DREAM registry did not contain data related to permanent work disability (PWD), we estimated time to PWD based on Chorus et al. [24]. The productivity cost due to PWD was estimated using the human capital approach based on the duration of PWD in patients with paid jobs, and average wages derived from Statistics Netherlands [33] for men and women in different age categories. The total productivity cost was calculated as the sum of costs due to sick leave and PWD.

Total drug cost for each patient was computed based on the simulated durations that the patient received specific drugs. Doses and unit prices of the drugs were determined based on documentation of the Dutch Institute for Health Care [37].

In the DREAM registry, HU was measured using the EuroQol five dimensions questionnaire (EQ-5D) [38, 39]. The histogram of the observed HU values showed bimodal normal distributions and therefore we performed a computer-assisted analysis of mixtures to divide the data into two latent classes [40]. Then, we used a two-part model to link HU to age, gender, disease duration, DAS28 and HAQ. In the first part, a logistic regression model was used to predict the probability that a patient belonged to a specific class. In the second part, a linear regression model, fitted for each specific class, was used to predict the HU (see Table 2). Predicted values of HU that fell outside of the realistic ranges were adjusted using the approach developed by Hernández Alava et al. [41]. In fitting each of the models for costs and HU, we performed backward variable selection using the AIC (see Table 2 for the final models).

QALYs were computed based on the predicted HU values at discrete time points. We selected QALY as an outcome because it is recommended for use as a standard measure of health effects in the Dutch and other guidelines for conducting and reporting economic evaluation [42-44]. Despite controversial assumptions underlying QALY, it has been by far the only generic measure of health that allows comparisons between diseases. Extensive discussion on the advantages and limitations of the QALY approach can be found in Loomes and McKenzie [45], Wagstaff [46], Broome [47] and Dolan [48]. Following the Dutch guidelines for pharmacoeconomic studies, annual discount rates for costs and QALYs were set at 4.0 and 1.5%, respectively [42].

2.3.4 Predicting the Health Assessment Questionnaire Disability Index

Because costs and QALYs are non-linearly related to DAS28 and HAQ (see previous section), we predicted HAQ based on disease duration, DAS28, age and gender using a linear mixed-effects model with random effects of intercept and disease duration. According to the rheumatologists' opinion, the long-term beneficial effect on HAQ of BRMs might be larger than that of DMARDs. Therefore, the models for HAQ progression in patients receiving DMARDs and BRMs were fitted separately using the Nijmegen Inception Cohort and DREAM registry, respectively. An overview of the model parameters, their estimates and distributions for probabilistic sensitivity analysis (PSA), and data sources for parameterisation is provided in Table 2.

2.4 Modelling Tools, Model Implementation and Output Analysis

The DES model was written using the Delphi language (Embarcadero Delphi XE 15.0; Embarcadero Technologies Inc., San Francisco, CA, USA). R [49] and SAS 9.2 (2008; SAS Institute Inc., Cary, NC, USA) were used for data handling and statistical analyses. The package CA.MAN was used for the mixture analysis of HU [50]. The algorithms provided by Ripley [51], Genz [52] and Press et al. [53] were used to generate random variables.
The model was rigorously verified for coding logics and correctness, and debugged based on extreme value scenarios.

Because RA is a chronic disease and the sequential treatment strategies included high numbers of drugs, a lifetime horizon was chosen to capture any differences in the long-term outcomes between the treatment strategies as recommended by the ISPOR Task Force on good research practices for modelling studies [54]. Therefore, the simulation was run until death of patients. The size of the initial cohort was determined by repeatedly running the simulation with increasing initial population size until means and standard deviations (SDs) of costs and QALYs became stable. As a result, an initial cohort of 10,000 patients was used for the first- and second-order uncertainty analyses. ICER was calculated according to Briggs et al. [55]. Ninety-five-percent confidence interval (CI) of ICER was computed using the non-parametric bootstrapping method [56] with 100,000 times of sampling. Simulated results were face-to-face validated by experts in RA and health economists.

For PSA, values of the model parameters were sampled 1,000 times from appropriate distributions. The beta distributions of the proportions (see Table 2) were estimated using the method of moments [55]. The multivariate normal distributions of the parameters in the regression models were obtained from the model fitting. The distribution of the logarithm of the scale parameters in the exponential distributions were estimated based on survival analyses [57]. The net-benefit framework was used to construct the cost-acceptability curves from the Monte Carlo simulation results [55].

3 Results

3.1 Disease Progression

The simulated results showed equal or better DAS28 and HAQ (as population averages) in Strategy 2 over time compared with Strategy 1 (Fig. 2). During the first 2 years, DAS28 and HAQ in both strategies decreased rapidly. After this period, however, HAQ in both strategies increased over time. The differences in HAQ between the two strategies were negligible during the first 2 years, increased gradually from year 2 to year 12, decreased gradually from year 12 to year 22, and became very small (<0.1) again after year 22.

3.2 Costs and Effectiveness

Mean societal costs per patient per year varied between €9,127 and €18,121 in Strategy 1, and between €10,731 and €18,027 in Strategy 2 over a period of 70 years. Overall mean (SD) of the non-drug direct cost per patient per year was €8,717 (2,937) in Strategy 1 and €8,548 (2,806) in Strategy 2, implying a marginally positive effect of BRMs on the non-drug resource use. Overall mean (SD) of the productivity costs per patient per year was €3,618 (9,101) in Strategy 1 and €3,379 (9,105) in Strategy 2. Overall mean (SD) of drug cost per patient per year was €155 (224) in Strategy 1 and €1,422 (2,242) in Strategy 2. Productivity and drug costs accounted for 48 and 4 % of the total costs in Strategies 1, respectively, and 34 and 31 % of the total costs in Strategy 2, respectively.

Fig. 2 Changes in means and standard deviations of DAS28 (a) and HAQ (b) of the patients in two treatment strategies over time. Strategy 1 consists of eight available DMARDs and Strategy 2 consists of the same available DMARDs as in Strategy 1 plus two anti-TNF and two non-TNF biological response modifiers. DAS28 Disease Activity Score for 28 joints, HAQ Health Questionnaire Assessment Disability Index, DMARDs disease-modifying antirheumatic drugs, TNF tumour necrosis factor
Fig. 3 Scatter plot of incremental mean costs against incremental mean QALYs of Strategy 2 compared with Strategy 1. Strategy 1 consists of eight available DMARDs and Strategy 2 consists of the same available DMARDs as in Strategy 1 plus two anti-TNF and two non-TNF BRMs. DMARDs disease-modifying antirheumatic drugs, QALYs quality-adjusted life-years, TNF tumour necrosis factor, BRMs biological response modifiers.

Fig. 4 Cost-effectiveness acceptability curves for two treatment strategies. Strategy 1 consists of eight available DMARDs and Strategy 2 consists of the same available DMARDs as in Strategy 1 plus two anti-TNF and two non-TNF BRMs. DMARDs disease-modifying antirheumatic drugs, TNF tumour necrosis factor, BRMs biological response modifiers.

In agreement with the trends of change in mean DAS28 and HAQ, mean HU in both strategies increased during the first 2 years and then steadily increased over time. The differences in mean HU between the two strategies were negligible during the first 4 years, increased from year 4 to year 12, decreased from year 12 to year 30, and became negligible (<0.01) again after year 30. Mean QALY per patient since RA diagnosis until death was 11.51 in Strategy 1 and 11.93 in Strategy 2 (95% CI of the difference 0.08–0.86). When combining costs and effects, ICER was €124,011. Bootstrapped 95% CI of ICER was (€62,193–597,857).

3.3 Probabilistic Sensitivity Analysis

The scatterplot of the joint uncertainty in the incremental mean costs against incremental mean QALYs of Strategy 2 compared with Strategy 1 showed that 23% of the data points lied in the north-west and remaining points lied in the north-east quadrants of the cost-effectiveness plane (Fig. 3). At a WTP threshold of €10,080, the probability that the Strategy 2 was cost effective was zero. At a WTP of €119,167, the probabilities that the Strategies 1 and 2 were cost effective were equal (0.5). At thresholds above €119,167, Strategy 2 cost-effectively dominated Strategy 1 but the probability that the Strategy 2 was cost effective never exceeded 0.87 (Fig. 4).

4 Discussion

As the treatment strategies for patients with early RA are becoming more and more complex, it is necessary to develop a modelling approach that fully supports economic evaluation of a wide range of real-life treatment strategies. While the ultimate goal of treating RA is to achieve remission, the desired treatment target may differ among countries. In our study, a clinical guideline was used to formulate the decisions when to change treatment and which drug should be used to maintain low disease activity (i.e. DAS28 ≤3.2) in RA patients. Low disease activity is accepted in international guidelines as a more realistic target compared with remission [58]. Using this treatment target, failure and switching of a treatment in the model were determined by DAS28, rheumatic factor and treatment duration. This approach is more realistic than using fixed sequences because it better reflects the practice of rheumatologists.

Owing to the flexible simulation approach, our model can be used to assess the effects of a number of factors on disease progression, costs and effectiveness such as: the number of BRMs available for a treatment strategy, the position of a specific drug in the sequential treatment, the cut-point of DAS28 as a criterion for switching treatment to reach low disease activity and the time interval between two assessments of disease activity. In addition, the model can be adapted to simulate the rheumatologist decisions on
drug use based on potential patient characteristics, of which values can be obtained, for example, using biomarkers. Although the model was developed for the current Dutch situation, it can be easily extended to incorporate different treatment algorithms based on new guidelines and guidelines from other healthcare systems.

In this study, we sought to estimate costs and effectiveness as they would occur in routine clinical practice. While data from controlled clinical trials are ideal for estimating treatment efficacy and for comparing outcomes between treatments in selective patients, they may not be relevant for economic evaluation studies in diverse patients under conditions different from those in the controlled trials [59]. Therefore, the use of real-life data for model inputs in our study is an advantage. Because the Nijmegen Inception Cohort and the DREAM registry were long-term follow-up studies that included patients with a wide range of disease severity, it was possible to link treatment effectiveness to DAS28 at the start of the treatment. In this way, the effectiveness is not overestimated in patients with mild RA or underestimated in patients with severe RA as in the case when the treatment effect is estimated independently of disease severity. Because the assumptions of normality and homoscedasticity of errors were satisfied for any ordinary least squares-based model fitted in our study, the predicted treatment effectiveness (using linear regression models) and costs (using log-linear regression models) were unbiased. We applied statistical models based on mixtures of parametric distributions for costs, which have been shown to perform better than those based on single distributions [60]. This approach was also applied for HU, of which the histogram clearly showed a mixture of two subpopulations. Recently, Hernández Alava et al. [41] have demonstrated that mixture models performed much better than standard linear regression models in predicting HU measured by the EQ-5D questionnaires. The simulated changes in mean DAS28 and HAQ in the two treatment strategies (Fig. 2) are plausible. During the first 2 years, most of the patients in Strategy 2 received DMARDs as those in Strategy 1. Therefore, means of these disease measures were almost identical. After 2 years, more and more patients in Strategy 2 were eligible to receive BRMs. Because the effectiveness of the first anti-TNF BRM was larger than that of DMARDs, and BRMs on average maintain DAS28 at low levels for a longer time than DMARDs, the numbers of patients having DAS28 returning to the baseline levels and of patients receiving 'palliative' therapy in Strategy 1 were increasingly higher than in Strategy 2, leading to the divergence in the differences between means of disease measures in the two strategies over the next 10 years. After 12 years, more and more patients in Strategy 2 received DMARDs after failures of all BRMs, narrowing the differences between the two strategies. After 25 years, a majority of the patients in the two strategies have lost response to all the available drugs and therefore the differences became negligible.

Our simulated ICER (£124,011) was higher than conventionally accepted WTP thresholds (e.g. £23,000 in UK and $62,000 in USA) [61]. In the models that evaluated treatment sequences independent of patient characteristics and used data from clinical trials for model parameterisation, estimated ICERs were lower than £100,000 [8, 9]. Inclusion of BRMs in a treatment strategy for RA appeared to be less favourable in our model, which used a real-life treatment algorithm and data. Analyses of the DREAM registry showed that the beneficial effect on DAS28 of BRMs was only slightly larger than those of DMARDs. Although in the long term, a patient treated with BRMs might have better physical functioning (i.e. lower HAQ) than treated with DMARDs (given the same baseline DAS28), this difference is rather small (see statistical models for HAQ in Table 2). These statistical analyses agree well with the simulated changes in DAS28 and HAQ over time in the two treatment strategies.

It should be noted that our estimated ICER was computed based on the expected costs and utilities of the whole population. Given the fact that HAQ progression, utilities and costs of patients with different ages, gender, disease duration and DAS28 are different, net benefit of a patient receiving a specific treatment clearly depends on his or her characteristics. As a consequence, there are always patients who do not benefit from the decided treatment strategy, leading to benefit forgone from the population-based decision. Thus, making a treatment decision for individual patients or subgroups of patients with similar characteristics would optimize the population welfare. Recently, much attention has been paid to an individual-based decision and the potential value that a society is willing to pay to realize the individualized care, the so-called expected value of individualized care (EVIC) [62, 63]. As our model took patient heterogeneity into consideration, it can be used to compute EVIC.

Our modelling approach has some limitations. It is complex and thus communication of the model structure and results with the decision makers may be not straightforward. In addition, extensive data are required for model parameterisation. Very large longitudinal studies are needed to quantify the effectiveness and toxicity profile of a drug at different positions and after stopping treatment of previous drugs because of different causes (primary failure, secondary failure or severe toxicity) in a treatment sequence. Because the contemporary management of RA aims at reaching low disease activity, the use of a high number of DMARDs is adopted. However, it currently remains uncertain about the effect on DAS28 of a DMARD after failure of several other DMARDs and BRMs. Because

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of insufficient data, in our model all DMARDs after failure of all BRMs were analysed as one class, and the effectiveness of the second anti-TNF and non-TNF BRMs were estimated independently of the cause of failure and toxicity of the first ones. Additionally, time to permanent work disability was modelled independently of the type of treatment. At the present time, it is still debatable whether or not treatment with BRMs could reduce the risk of becoming permanently work disabled [94]. Nevertheless, while complexity and demand for data are limitations of our model, they are helpful in identifying the knowledge gaps in the current guidelines for the management of RA and priorities for future research.

5 Conclusions

It is possible to model the outcomes of complex treatment strategies based on a clinical guideline for RA. This modelling approach allows assessment of the impact of changes in clinical practice on disease progression, health and costs of the RA patients, and cost effectiveness of new treatment decisions. Following the Dutch guideline for the management of RA and using real-life data, the simulated ICER of the treatment strategy including BMRs compared with a treatment strategy without BMRs was higher than conventionally accepted WTP thresholds. Using an individual-oriented approach, our model can serve as a tool for individualized care analysis taking patient heterogeneity into account and can be extended to assess total health and budget impact of treatment strategies in different decision-making contexts. Despite complexity and demand for extensive data, our modelling approach can help to identify the knowledge gaps in clinical guidelines for RA management and priorities for future research.

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Author contributions  ATD designed and programmed the model, reviewed literature and performed statistical analyses for model inputs, ran the simulations, debugged the model, analysed the model outputs, designed and created the figures, and drafted the manuscript. AB assisted in drafting the manuscript. WK compiled the input data and estimated costs and utility. AB, WK, PLCMVr and MAFIvDL provided the first algorithm for treatment decisions. JLS assisted in conceptualising the model. All authors participated in discussion and refinement of the algorithm for treatment decisions, interpretation of the simulated results and review of the manuscript. ATD will serve as a guarantor for the entire contents of the manuscript.

References


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